Banking and Interest Rates in Monetary Policy Analysis:

A Quantitative Exploration

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1. Introduction

Recent years have seen great changes in monetary policy analysis, as economists in central banks and academia have come together on an analytical approach of the general type discussed by Rotemberg and Woodford (1997), Goodfriend and King (1997), Clarida, Galí, and Gertler (1999), Woodford (2003), and many others. This approach is characterized, as argued by McCallum (1999), by investigations of alternative rules for monetary policy conducted in models that are based on private-agent optimizing behavior but with specifications that include features designed to lend empirical veracity, thereby aspiring to be structural and accordingly usable (in principle) for policy analysis. Despite a widespread belief that this approach is fundamentally sound, and that recent work represents a major improvement over the practice typical 15 or 20 years ago, there are some reasons for unease. Prominent among these are the absence from the standard framework of any significant role for monetary aggregates, financial intermediation, or distinctions among various short-term interest rates that play different roles in the transmission mechanism.

A recent paper by Goodfriend (2005) develops a qualitative framework designed to overcome these particular weaknesses. Specifically, it “… integrates broad money demand, loan production, asset pricing, and arbitrage between banking and asset markets” (2005, p. 277) and illustrates the logical necessity (in principle) for monetary policy to take account of—among other things—the difference between the interbank rate of interest (used as the policy instrument) and other short rates including the government bill rate, the collateralized bank loan rate, the (nominal) net marginal product of capital, and a shadow (total) nominal risk-free rate—each of which differs from the others. As noted by Hess (2005), however, Goodfriend (2005) provides no evidence or argument concerning the quantitative importance of these features and distinctions. The primary objective of the present paper, accordingly, is to formulate a quantitative version of Goodfriend’s model, develop a plausible calibration, and utilize this model to assess the magnitude and policy
relevance of the effects and distinctions just mentioned for steady state interest rates and aggregate variables, and for dynamic monetary policy simulations. Among other things, the paper will investigate the role of the “external finance premium” that is emphasized in the prominent work of Bernanke, Gertler, and Gilchrist (1999). It will do so using a model in which the external finance premium is endogenously determined by no-arbitrage relationships in an environment in which loan production depends upon both collateral and loan-monitoring inputs, with capital serving less efficiently as collateral than bonds, while medium-of-exchange money is crucial for facilitating transactions. In this setting, unlike that of Bernanke, Gertler, and Gilchrist (1999), the external finance premium may move either pro-cyclically or counter-cyclically in response to shocks, depending upon parameters of the model.

How does the present paper compare with previous efforts to outline and quantify the role of financial intermediation (banking) in monetary policy? Probably the most prominent line of work of this type is that begun by Bernanke and Gertler (1989, 1995) and continued by Bernanke, Gertler, and Gilchrist (1999), but the literature also includes notable contributions by Kiyotaki and Moore (1997), Carlstrom and Fuerst (1997), Kocherlakota (2000), Cooley, Marimon, and Quadrini (2004), and others. It is apparently the case, however, that in all of these studies the models are fundamentally non-monetary—i.e., do not recognize the existence of a demand for money that serves to facilitate transactions.¹ This omission could be of first-order importance for the financial accelerator, however, for its mechanism works via increases in the supply of collateral induced by asset price increases. In models with money, however, such increases also increase the demand for collateral as spenders go to the banking system for additional money to facilitate the additional spending induced by the initiating shock. Accordingly, our analysis focuses on the net effect of these offsetting forces.

¹ Diamond and Rajan (2006) is a noteworthy qualitative study of the role of banking in monetary policy.
Our model’s “banking accelerator” transmission channel works in much the same way as the financial accelerator does in existing models. For instance, monetary policy that stimulates employment and output in the presence of sticky prices raises the marginal product of capital, the price of capital, and the value of collateral in the economy, thereby tending to reduce the external finance premium for a given quantity of bank deposits demanded. However, our model includes in addition a “banking attenuator” channel, which recognizes that monetary stimulus to spending also increases the demand for bank deposits, thereby tending to raise the external finance premium for a given value of collateral-eligible assets in the economy.

Also of some indirect relevance to the present exploration are the studies of Ireland (2003), McCallum (2001), and Woodford (2003, pp. 300-311), which seek to quantify the effects of neglecting monetary aggregates in specifications that do not adopt the usual approximation that is necessary to keep monetary aggregate magnitudes from appearing in the intertemporal optimizing conditions of a standard policy-analysis model. These papers, all of which found only small effects of this approximation, did not include any explicit banking sector, however, and did not recognize distinctions among various short-term interest rates.

In effect, ours is a two-sector model with a goods-producing sector and a banking sector. Goods are produced with capital and work effort as usual. The banking sector produces loans and deposits according to a production function with inputs of monitoring effort and collateral, the latter consisting of government bonds (held by the private sector) and capital employed in the goods-producing sector. The distinction between various interest rates arises in the model because loans and deposits are costly to produce in the sense that they require work effort, while collateral services allow an economization of that effort. Hence, the total return on bonds and capital has an explicit pecuniary component, and an implicit liquidity services yield component that reflects the collateral services that these assets provide. Portfolio balance requires risk-adjusted equality
between the various assets’ total returns. Thus the sum of pecuniary and service-yield returns must
equal a shadow total nominal risk-free interest rate. However, the pecuniary bond rate is less than
the net nominal marginal product of capital because bonds are more productive as collateral than
capital. In equilibrium the interbank rate, which is the cost of loanable funds for a bank, is below the
(uncollateralized) loan rate by the marginal cost of loan production. Finally, the loan market and the
asset markets are linked by a no-arbitrage condition between the uncollateralized loan rate and the
shadow total nominal risk-free rate.²

The strategy of the paper is two-fold. First, it is to use observed average historical values of
interest rates and rate spreads, together with observations on banking and macroeconomic
aggregates, to calibrate steady-state equilibrium values for the model with realistic trend
productivity growth in the production of goods and loans. Our aim in this regard is to determine the
extent to which the introduction of money and banking into an otherwise standard growth model
can account for observable interest rate differentials, and how much money and banking matters
quantitatively (on average) for aggregates like the capital stock, employment, and output. In this
regard, we develop the implications of money and banking with reference to the famous equity
premium puzzle of Mehra and Prescott (1985), summarized by Campbell (1999) as the puzzle of
“why the average real stock return is so high in relation to the short-term interest rate.”

The second part of our strategy is to log-linearize the model around the calibrated steady
state to explore how much the inclusion of money and banking matters quantitatively for an
otherwise standard “new neoclassical synthesis” (aka, “new Keynesian”) model of monetary policy.
In particular, we wish to investigate quantitatively how much a central bank can be misled by
relying on a new neoclassical synthesis model without money and banking when managing its
interbank rate policy instrument.

² Different financial assets play an important role in the model’s transmission mechanism, in a manner that is somewhat
We begin by illustrating the presence of the two channels—the accelerator and the attenuator—by which money and banking influence the model economy’s dynamics. Next, we illustrate the extent to which a central bank could misjudge its interbank rate response to a goods productivity shock by not taking money and banking into account. Finally, we consider shocks emanating from the banking sector itself—a shock to loan monitoring productivity and a shock to effective collateral that is meant to capture widespread financial distress.

The paper’s outline is as follows. In Section 2, we begin by reviewing and building upon the specification of the Goodfriend (2005) model and highlighting some of its features. Next, in Section 3 we emphasize the various interest rates that appear in the model and the frictions that make them differ from each other. Then in Sections 4, 5, and 6 we develop the steady state solution that forms the basis for the linearized version of the model, which will be used in subsequent analysis, and discuss the steady state calibration and some of the quantitative consequences of incorporating money and banking. In Section 7 we complete the specification and linearization of our dynamic model. Finally, in Section 8 we conduct various policy experiments to see how this model performs in comparison to more standard specifications without any banking sector. Conclusions are briefly outlined in Section 9.

2. Core Model Outline

The model that we will use to investigate these issues regarding the role of money and banking needs to be specified at the level of preferences and technology, so as to be potentially structural and invariant to policy specification. In order to keep the analysis as simple as possible, we have specified the model in terms of an optimizing problem for a typical household, which not only consumes a bundle of differentiated commodities, supplies labor, and saves but also produces for sale a differentiated good and in addition operates a (competitive) banking firm. Thus the basic optimizing analysis is conducted in one step. It will take a bit of discussion, nevertheless, to
describe the details of the typical household’s decision problem. When that task is completed, we will specify that there are many households and aggregate by assuming that they are all similar and finding a symmetric equilibrium.

The typical household’s objective function is to maximize, at time $t = 0$, the value of

$$E_0 \sum_{t=0}^{\infty} \beta^t [\phi \log(c_t) + (1 - \phi) \log(1 - n_t^s - m_t^s)]$$

where $c_t$ is consumption during period $t$ of Dixit-Stiglitz consumption bundles and $(1 - n_t^s - m_t^s)$ is leisure in $t$, with $n_t^s$ representing labor supplied to the production of goods (in the household firm or elsewhere) and $m_t^s$ is labor likewise supplied to banking activities (to be described below). In the household’s budget constraint, we must recognize that the household pays for its consumption expenditures with income from its production and sale of goods, net receipts from supplying labor, net sales of its financial assets (money and bills$^3$), and net sales of capital goods. (For simplicity, we are going to keep the aggregate stock of capital constant—at a steady-state magnitude that is determined endogenously—but individual households can buy or sell capital goods so their market price is a variable.) The household has market power in the product that it supplies, with $\theta$ being its elasticity of demand, stemming from a value of $\theta$ for the elasticity of substitution in consumption in the usual Dixit-Stiglitz aggregator setup. Consequently, the household maximizes (1) subject to two constraints. The first is the budget constraint

$$q_t (1 - \delta)K_t + \frac{B_t}{P_t^A} + \frac{H_{t-1}}{P_t^A} + w_t (n_t^s + m_t^s) + c_t^A \left( \frac{P_{t+1}}{P_t^A} \right)^{1 - \theta}$$

$$- w_t (n_t^s + m_t^s) - \frac{H_t}{P_t^A} - tax_t - q_t K_{t+1} = \frac{B_{t+1}}{P_t^A (1 + R_t^B)} - c_t = 0,$$

and the second is the sales equals net production constraint needed to reflect monopolistic

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$^3$ Bills are one-period government bonds. Note that asset-liability positions between households and banks are netted out, a step that requires that banking firms possess no monopoly power.
competition in the supply of the household’s own produced good, namely,
\[ K_t^n (A_t n_t)^{1-n} - c_t^\Lambda (P_t / P_t^\Lambda)^{-\theta} = 0. \]

Here the variables are as follows: \( q_t \) = real price of capital, \( K_t \) = quantity of capital at start of \( t \), \( B_{t+1} \) = nominal bills held at end of \( t \), \( P_t \) = price of household’s produced good in \( t \), \( P_t^\Lambda \) = consumption goods price index, \( n_t \) = labor demanded by household as producer, \( m_t \) = labor demanded by household’s banking operation, \( w_t \) = real wage rate, \( H_t \) = nominal holdings of base money held at the end of \( t \), \( \text{tax}_t \) = real lump-sum tax payments in \( t \), and \( R_t^B \) = nominal interest rate on government bills. The Lagrangian multiplier on (2) is denoted \( \lambda_t \) and for (3) it is \( \xi_t \), with \( \beta_{t-1} \) inserted (in both cases) in the Lagrangian expression for each period to make the multipliers represent values as of the indicated period \( t \), not the startup date 0. In (3), \( A_1 \) is a shock to productivity in goods production, similar to a standard technology shock in the real-business-cycle literature, whose mean increases over time at the trend growth rate \( \gamma \). The superscript \( \Lambda \) on any variable, it might be noted, indicates that the variable is an economy-wide measure that each individual household takes as given.

In addition, the model includes one more constraint, which will build in most of our money-and-banking content. Since we require that the model must pertain to a monetary economy, we need to incorporate the transaction-facilitating properties of money—the medium of exchange—in some explicit manner. The most concrete way possible is via a constraint requiring the household to pay for consumption spending during period \( t \) with money held in that period. Our specification will have the medium of exchange composed entirely of bank deposits, and will require that consumption during \( t \) must be rigidly related to deposits held at the end of \( t \). We assume that this constraint always binds, and as a consequence will require below that \( c_t = V D_t / P_t^\Lambda \), where \( D_t \) is

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4 The relative price in terms of consumption goods.
nominal deposits and V is the constant that reflects the assumed rigidity.\(^5\)

In this context, we come to the banking sector. We write the simplest possible bank balance sheet, namely, \(H_t + L_t = D_t\), where \(H\), \(L\), and \(D\) are high-powered (base) money, loans to households, and deposits, respectively, for the typical bank.\(^6\) We let \(rr\) be the bank’s chosen ratio of \(H_t\) (equals bank reserves) to deposits, and from \(rrD_t + L_t = D_t\) obtain \(D_t = L_t/(1 - rr)\), so that we have the transaction constraint

\[
(4) \quad c_t = \frac{VL_t}{(1 - rr)P_t^\alpha}
\]

and also an identity

\[
(5) \quad H_t = \frac{(rr/(1 - rr))L_t},
\]

which represents a derived demand for high-powered (base) money. The analytical heart of the banking sector specification is a model of loan “production” or, more accurately, loan management. This crucial activity is conducted, we assume, by a combination of collateral requirements and loan monitoring, the latter activity being performed by workers whose labor is supplied by households to the banking sector, as mentioned above. Specifically, we posit the following production function pertaining to the management of loans:

\[
(6) \quad L_t / P_t^\alpha = F(b_{t+1} + A3, kq_t K_{t+1})^\alpha (A2, m_t)^{1-\alpha}.
\]

Here we have a Cobb-Douglas production function with factor inputs being labor \(m_t\) (for monitoring) and collateral \(b_{t+1} + A3, kq_t K_{t+1}\), with \(b_{t+1} = B_{t+1} / P_t^\alpha (1 + R_t^B)\).\(^7\) In the latter expression the effective amount of collateral possessed by a typical household depends upon, among other

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\(^5\) This is not our favorite specification; we plan to loosen it in subsequent work. Given, however, that we are using this constraint—similar in spirit to a “cash in advance” constraint—it is necessary to incorporate some factor such as V to take account of the fact that the velocity of aggregate bank deposits is far less than 1.0 for quarterly time periods. We follow the usual convention of calibrating the model so that time periods represent quarters.

\(^6\) Kashyap, Rajan, and Stein (2002) offer a microeconomic explanation of the synergy between deposit and loan provision in banking.

\(^7\) Banks expend effort and require borrowers to post collateral, but in equilibrium there is no default. This modeling choice is based, in part, on progress that has been made in understanding the implications of credit market imperfections in limited commitment environments where there is no equilibrium default, such as Kocherlakota (1996).
things, the price of capital goods, \( q_t \). The constant \( k, 0 < k < 1 \), reflects the inferiority of capital to bonds for collateral purposes, resulting because capital goods (unlike bonds) typically require substantial monitoring to verify their physical condition and market value. The variables \( A_2 \) and \( A_3 \) are shock terms, reflecting stochastic technology effects, analogous to \( A_1 \) in the goods production function included in (3). Here \( A_2 \), like \( A_1 \), has a trend growth rate of \( \gamma \) whereas \( A_3 \) is trendless.

With the model as specified, we derive first-order optimality conditions pertaining to the choice variables \( m^*, m, n^*, n, K, P, \) and \( B^* \), and also define the variable  

\[
\Omega_t = \alpha c_t / (b_{t+1} + kq_tK_{t+1}).
\]

Then imposing symmetry and market clearing, and also constant aggregate capital, we obtain the following conditions for equilibrium:

\[
\frac{1 - \phi}{1 - n_t - m_t} = w_t \lambda_t 
\]

\[
\frac{w_t m_t}{c_t (1 - \alpha)} = \frac{\phi}{c_t \lambda_t} - 1 
\]

\[
w_t = (\xi_t / \lambda_t) A_1 (1 - \eta) \left( \frac{K}{n_t A_1} \right)^\eta 
\]

\[
\left( \frac{\phi}{c_t \lambda_t} - 1 \right) k \Omega_t q_t - q_t + \beta (1 - \delta) E_t \left( \frac{\lambda_{t+1}}{\lambda_t} q_{t+1} \right) + \beta \eta E_t \left( \frac{\lambda_{t+1} \xi_{t+1}}{\lambda_t \lambda_{t+1}} \left( \frac{A_{1t+1} n_{t+1}}{K} \right) \right) = 0
\]

\[
\left( \frac{\phi}{c_t \lambda_t} - 1 \right) \Omega_t - 1 + \beta E_t \left( \frac{\lambda_{t+1} P_t}{\lambda_t P_{t+1}} \left( 1 + R^B_t \right) \right) = 0
\]

\[
\xi_t / \lambda_t = (\theta - 1) / \theta 
\]

\[
b_{t+1} = B_{t+1} / P_t^\lambda (1 + R^B_t) 
\]

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8 The system has some recursivity such that \( c_t \) and \( H_t \) can be solved for after the other variables.

9 The role of variable \( \Omega_t \) will be described in Section 3.
Here (15) is the government (including central bank) budget constraint. In a symmetric, flexible-price equilibrium, the household budget condition [implied by (2) and (3)] plus equations (4)-(15) serve to determine values of the thirteen endogenous variables $c, n, m, w, q, \lambda, \xi, L, \Omega, B, \text{tax}$, and $R^B$ given exogenous shock processes and policy-set time paths for $H, g,$ and $b$. Alternatively, and more realistically, we will also study equilibria in which the central bank sets values of an overnight interest rate, rather than $H_t$. In the foregoing list of variables, we have not included the capital stock, $K$, because our equilibria are defined with $K_t$ equal to its steady-state value in each period. The steady-state value of $K$ is, however, determined endogenously.

In our experiments involving dynamic adjustments we will, however, use a sticky-price version of the model by eliminating the labor demand equation (10) and replacing it with a standard Calvo price adjustment relation, in a manner that will be described below in Section 5.

3. Interest Rates

Now we turn our attention to the various interest rates mentioned in the introduction. With the exception of the bill rate for loans to the government, these did not appear in the analysis of Section 2 because of our expositional strategy of treating goods producers and banks as owned and operated by households. Here, we introduce a one-period default-free bond that provides no collateral services to its holder. We denote the nominal rate on this fictitious security as $R^T_t$, as in Goodfriend (2005), and will use it as a “benchmark” rate in what follows. From the specification of the household optimization problem above, it is clear that $R^T_t$ satisfies the familiar condition

\begin{equation}
1 + R^T_t = E_t \frac{\lambda_t}{\beta \lambda_{t+1}} \frac{P_{t+1}}{P_t}.
\end{equation}

How does this rate relate to the bill rate that appeared above, $R^B_t$? By comparison with first-order

\footnote{For simplicity, we will keep government spending $g$, equal to zero throughout.}
condition (12) for optimal household holdings of bills, we see that

$$\frac{1 + R^B_t}{1 + R^T_t} = 1 - \left( \frac{\phi}{c_i \lambda_i} - 1 \right) \Omega_t.$$ 

Thus these two rates coincide if, and only if, $\Omega_t = 0$. That property makes good intuitive sense, as $\Omega_t \geq 0$ can be viewed as the partial derivative of the constraint (4) with respect to collateral, and therefore the rates differ only if the bills are collateralizable. Furthermore, $\left( \frac{\phi}{c_i \lambda_i} - 1 \right) \geq 0$ with the strict inequality holding whenever (12) is binding. Therefore we can identify $\left( \frac{\phi}{c_i \lambda_i} - 1 \right) \Omega_t$ as the liquidity service yield on bonds, which we denote as $LSY^B_t$. From (17), then, we have the approximate equality

$$R^T_t - R^B_t = LSY^B_t.$$ 

In our model’s non-stochastic steady state, bills and capital will have the same risk properties, but capital will serve as collateral less well than bills, by the factor $k$. Accordingly, one might guess that the liquidity service yield on capital, $LSY^K_t = R^T_t - R^K_t$, would satisfy the approximation

$$LSY^K_t = kLSY^B_t.$$ 

In fact, that result can be verified by a more complex calculation analogous to the one conducted for bills.\(^{11}\)

A crucial interest rate for monetary policy analysis is the interbank rate, denoted $R^IB_t$. Since such a rate serves as the policy rate for the Federal Reserve and many other central banks, $R^IB_t$ will be viewed as the policy instrument in some of our experiments in Section 8. How, then, is the

\(^{11}\)This result, and the one in the previous paragraph, is more fully developed in Goodfriend (2005).
interbank rate \( R^\text{IB} \) related to other rates? The basic fact is that, in our model, a bank can obtain funds from the interbank market at the rate \( R^\text{IB} \) and could loan them to households, if it incurred all the necessary costs, at the benchmark rate \( R^\text{T} \). This reflects a no-arbitrage condition between the loan market and the asset market. Then what are these costs? We have postulated above that loan production requires monitoring costs and collateral costs as specified in equation (6). Here we temporarily pretend that the collateral costs are borne by the bank, i.e., we consider uncollateralized loans. At the cost-minimizing optimum mix of factor inputs, the real marginal cost of loan production equals the factor price divided by the factor’s marginal product for all factors of production. Thus we can find this marginal cost as the real wage divided by the partial derivative of \( L_t/P_t \) with respect to \( m_t \), where the latter is employment of labor services (for monitoring) by the typical bank. From (6) we find this partial derivative to be \( (1 - \alpha)(L_t/P_t)/m_t \) so using expression (4) we obtain \( V m_t w_t/(1-\alpha)(1-rr)c_t \) as the real marginal cost of loan management. Profit maximization by banks implies, then, that

\[
(1 + R^\text{IB}) \left[ 1 + \frac{V w_t m_t}{(1-\alpha)(1-rr)c_t} \right] = (1 + R^\text{T}),
\]

which can be approximated as

\[
R^\text{T} - R^\text{IB} = \left[ \frac{V w_t m_t}{(1-\alpha)(1-rr)c_t} \right].
\]

Now, although we can imagine uncollateralized loans to households at the “total” asset market rate \( R^\text{T} \), actual loans are collateralized in equilibrium. Accordingly, since \( (1 - \alpha) \) is the factor share for monitoring, the marginal cost of loan production is multiplied by \( (1 - \alpha) \) and the relevant relationship becomes
\[(22) \quad (1 + R_i^{\text{in}}) \left[1 + \frac{V_{w,m_i}}{(1-\tau r)c_i}\right] = (1 + R_i^{\text{l}}),\]

where \(R_i^{\text{l}}\) is the interest rate on collateralized loans. As before, the corresponding approximation is

\[(23) \quad R_i^{\text{l}} - R_i^{\text{in}} = \left[\frac{V_{w,m_i}}{(1-\tau r)c_i}\right].\]

Finally, in our setup, banks are paying households a rate \(R_i^{\text{D}}\) on their deposits. Given competitive conditions, this rate differs from the interbank rate only because the fraction \(1-\tau r\) of non-interest bearing deposits are not loaned, implying that

\[(24) \quad R_i^{\text{D}} = R_i^{\text{in}}/(1-\tau r).\]

Given this plethora of interest rates and rate differentials, which differential is it that measures the “external finance premium” in the context of our model? We identify the external finance premium with the real marginal cost of loan production, since our loan production costs reflect the cost of external finance emphasized among others by Bernanke, Gertler, and Gilchrist (1999). Household borrowers pay a loan rate that covers real marginal cost in addition to the interbank rate, which equals the deposit rate except for a small reserve ratio. Hence, the real marginal cost of loan production is an external finance premium from the household’s perspective.

It is useful to distinguish between an uncollateralized and a collateralized external finance premium in our model. The external finance premium on an uncollateralized loan (UEFP) would be \(R_i^{\text{T}} - R_i^{\text{in}}\) because this interest rate spread reflects the full real marginal cost of loan production. In effect, \(R_i^{\text{T}}\) equals the uncollateralized loan rate.

On the other hand, the spread between the collateralized loan rate and the interbank rate is \(R_i^{\text{l}} - R_i^{\text{in}} = (1-\alpha)(R_i^{\text{T}} - R_i^{\text{in}})\) because the smaller \(R_i^{\text{l}} - R_i^{\text{in}}\) interest rate spread need only cover the portion of real marginal loan cost due to the monitoring effort required when the borrower puts
up the requisite collateral. In effect, households that borrow on a collateralized basis receive a
deduction on the loan rate equal to the share $\alpha$ of collateral in loan costs. For calibration, we
identify the collateralized external finance premium, which we denote as CEFP, with the spread
between the prime rate and the interbank rate.

Because the banking sector is perfectly competitive and loan production features constant
returns to scale, net interest income $(R_t^T - R_i^M) L_t/P_t$ equals the sum of payments to the factor
inputs in loan production, $m_t w_t + b_{t+1} LSY_t^B + q_t(K_{t+1}) LSY_t^K$. The zero profit condition indicates
that the liquidity service yields on bonds and capital can be interpreted as “rent” paid to borrowers
for the collateral services that their assets provide in loan production. Bonds and capital earn a
lower pecuniary return than $R_t^T$, the fictitious benchmark bond that provides no collateral services,
by an amount equal to the “rent” that bonds and capital earn as a deduction on borrowing costs.

The external finance premium and the liquidity service yields on bonds and capital are
related to each other as $LSY_t^K = k LSY_t^B = k \Omega_t [(1-\alpha)/V(1 - \alpha)] CEFP_t$.

4. Steady State Solution

The next step is to present the deterministic, zero-inflation, steady-state version of the model
in Section 2 as a basis for calibration and our first batch of analytical exercises. Preliminary
consideration of the steady-state system indicated that the inclusion of output growth would be
essential to any degree whatsoever of realism. Accordingly, we assume that the shock terms $A_1$ and
$A_2$ in the production functions for goods and loans each trend upward at the growth rate $\gamma$.
Therefore deterministic expressions for these two variables can be written as $A_1 = A_{10}(1 + \gamma)^t$ and
$A_2 = A_{20}(1 + \gamma)^t$. For simplicity we normalize by setting $A_{10} = A_{20} = 1$. For a steady state, $\lambda_t$ must
shrink at the rate $\gamma$. Next we note that the relative price of capital will, in the steady state, equal 1,
so that detrended average per-household capital $K_t$ is an endogenous constant, K. Consequently, by substituting out certain variables we can express the core steady-state system in terms of seven equations involving only the seven variables $c, m, n, w, \lambda, \Omega,$ and $K$.\textsuperscript{13} Here $c, w, \lambda,$ and $K$, are constant detrended values of the variables written above, which grow (or shrink) at the rate $\gamma$, while the others are constant without any detrending. We will include $q$ in the expressions to reveal its importance in the system, but the reader should keep in mind that $q = 1$ and that $k$ is a parameter. As a piece of notation, we use $boc$ to denote the constant steady-state value of bonds to consumption and assume that the fiscal authority behaves so as to stabilize $boc$ around an exogenous, policy-determined value.

The first step is to substitute (6) into (4) and write the result as

\begin{equation}
1 = \frac{VF}{(1-\eta)} \left( boc + \frac{kqK}{c} \right)^{\frac{\alpha}{\gamma}} \left( \frac{m}{c} \right)^{1-\alpha}.
\end{equation}

Then equations (7) – (11) above can be represented as follows:

\begin{align}
(26) \quad \Omega &= \frac{\alpha}{boc + kqK / c} \\
(27) \quad \frac{1 - \phi}{1 - n - m} &= w\lambda \\
(28) \quad \frac{wm}{c(1-\alpha)} &= \left( \frac{\phi}{\lambda c} - 1 \right) \\
(29) \quad w &= \frac{(\theta - 1)(1-\eta)}{\theta} \left( \frac{K}{n} \right)^{\eta} \\
(30) \quad \left( \frac{\phi}{c\lambda} - 1 \right) k\Omega - 1 + \frac{\beta}{1 + \gamma} \left( 1 - \delta + \frac{\eta(\theta - 1)}{\theta} (K/n)^{\eta-1} \right) &= 0
\end{align}

Finally, we also have one relation that amounts to the overall resource constraint that incorporates

\textsuperscript{13} In addition, the variable $r$, which denotes the before-depreciation marginal product of capital, could be obtained recursively after the other seven variables are determined from $r - \delta = (\eta(\theta-1)/\theta)(n/K)^{1-\eta}$. 

15
equations (2), (3), and (15); it is as follows:

\[(31) \quad 1 = (K / c)^n (n / c)^{1-n} - \delta K / c.\]

Thus we have the desired seven equations, (25)-(31), to determine the steady-state values of c, m, n, w, \(\lambda\), \(\Omega\), and K, and these are the only included variables since boc and q are exogenous.

Given these core-model variables, and a given steady-state inflation rate, one can then readily solve for steady-state values of \(R^T\), \(R^B\), \(R^B\), \(R^L\), and \(R^D\) using equations (16), (17), (21), (23), and (24). Accordingly, we now turn to our calibration.

### 5. Steady-State Calibration

One leading objective in the paper is to explore the quantitative consequences of money and banking for interest rates and monetary policy in an otherwise conventional benchmark new neoclassical synthesis (new Keynesian) model. Because the model in this paper is not as fully specified as we would like, we have thus far not attempted to develop a full-fledged, precise calibration. Our aims are more modest. First, we want to determine whether a steady-state calibration exists that fits the relevant observed variables tolerably well. Second, we want to explore under this calibration how much money and banking matter quantitatively for the various interest rates and the model’s other endogenous macroeconomic variables.

The model has twelve parameters that need to be specified. Six of the parameters appear in the benchmark model without money and banking. For these we follow convention. First, we calibrate \(\phi = 0.4\) to yield roughly 1/3 of available time working in either goods production or banking. Second, we choose \(\eta = 0.36\) to reflect relative shares of capital and labor in goods production. Third, we choose \(\theta\) to yield a markup of 1.1. Fourth, we choose a conventional quarterly \(\beta = 0.99\). Fifth, we set the quarterly depreciation rate \(\delta = 0.025\). Sixth, we set \(\gamma = 0.0050\) to correspond to 2% productivity growth per year.

We calibrate the fiscal policy parameter boc by using values of U.S. public debt held
domestically by private entities, relative to GDP, as of the third quarter of 2005. The resulting figure is boc = 0.56.

That leaves five money and banking parameters to calibrate. First, we calibrate the velocity of aggregate bank deposits as the ratio of GDP to M3 for the fourth quarter of 2005 so that \( V = 0.31 \). Second, we set the reserve ratio at \( r_r = 0.01 \), measured as the ratio of total bank reserves to M3 as of December 2005. Finally, we calibrate the remaining three banking parameters \( \alpha \), \( k \), and \( F \) jointly to match, as best as we can, the following conditions: (i) a 1% per year average short-term real “riskless rate” that is the benchmark in the finance literature, e.g. Campbell (1999, p. 1233), (ii) a 2% average external finance premium that is in line with the average spread of the prime rate over the federal funds rate in the postwar United States, and (iii) a share of total US employment in depository credit intermediation as of August 2005 of 1.6% as reported by the Bureau of Labor Statistics. Accordingly, the three remaining parameters are calibrated as \( \alpha = 0.65 \), \( k = 0.2 \), and \( F = 9.0 \).

Our model of money, banking, and interest rate spreads was constructed on what we regard as a sensible theoretical foundation. It was not at all obvious to us that the banking parameters in the context of the macro model could achieve a reasonable calibration. Nevertheless, we shall see below that the calibration works tolerably well and in our view does provide a basis for judging the quantitative importance of money and banking in the steady state and in the dynamic analysis of monetary policy.

**6. Steady-State Analysis**

We begin the analysis of the steady state below by describing how well the benchmark calibration given in Section 5 is able to match the endogenous interest rates and aggregates to the respective observables. Then we illuminate the quantitative consequences of money and banking for the steady state in two ways. First, we compute a counterfactual calibration by presuming that the
banking sector is 10 times more efficient in loan production than in the benchmark calibration. We compare the “highly efficient banking” and the benchmark calibrations to show how much money and banking matters quantitatively. Second, we perform two experiments on the benchmark calibration to illustrate its sensitivity to the debt to GDP ratio (boc), the velocity of aggregate bank deposits (V).

The steady-state values of the endogenous variables for the benchmark calibration are reported in the top panel of Table 1. Starting with the labor market variables, we see that total available working time is close to 1/3 as desired. Moreover, the ratio of time worked in banking to total work effort, m/(n+m), is 1.8% under the benchmark calibration, not too far from the 1.6% ratio in the US economy. The steady-state capital/output ratio is 2.7 annualized, which is also in an acceptable range.

Turning to the interest rates, note that the steady state is computed at zero inflation so that we can interpret all the interest rates as “real” rates. To start, note that the rate of interest R^T on the benchmark fictitious security that provides no collateral services is exactly 6% per annum. This follows directly from the calibrated values of β and γ, together with log utility, which imply that R^T = ρ + γ, where β = 1/(1 + ρ). Strictly speaking, the interbank rate R^IB and the government bill rate R^B are both short-term interest rates in the model, so they should both be near the 1% per annum average short-term rate. The calibration puts the interbank rate at 0.8% per year, which is tolerably close to 1%. However, the government bill rate is 2% per year under the calibration, which is less satisfactory. That said, the two rates do straddle the 1% per annum average short-term rate, and neither is too far from 1%. We believe that replacing the Cobb-Douglas loan production function with a CES production function might help with this problem.

It is noteworthy that the average maturity of the stock of public debt outstanding used to calibrate the model is around 5 years. On this basis one might presume that R^B under the calibration
is more representative of a 5-year bond rate than the short-term bill rate. In this regard, it is interesting to note that Campbell, Lo, and MacKinlay (1997, p.415) report an average 5-year term premium of around 1.2% per year relative to the short rate, which happens to match closely the spread between $R^B$ and $R^B$ in our calibrated steady state. Such evidence provides a plausible reason why $R^B$ is elevated under our calibration, and suggests that it would be useful to modify the model to distinguish various maturities of government bonds outstanding in order to improve the calibration.

The 2% per annum external finance premium under the calibration lines up well with the average spread between the prime rate and the interbank rate. This so because the banking parameters were chosen to approximate closely the spread on collateralized loans, $R^l - R^B$.

In order to illuminate the extent to which money and banking matter quantitatively for steady-state interest rates and macroeconomic aggregates, consider a counterfactual calibration in which loan production is 10 times more efficient than in the benchmark calibration. Specifically, suppose that $F = 90$. The endogenous steady-state variables produced by the counterfactual calibration are shown in the bottom panel of Table 1 labeled “highly efficient banking.” The striking feature of the “highly efficient banking” steady state is that all the interest rates have converged to $R^T$, and the external finance premium and work in banking have shrunk so much that Matlab reports them both as approximately zero.

We can interpret the counterfactual as a steady state in which banking services are essentially free and compare it to the benchmark steady state in order to compute how much money and banking matter quantitatively for the endogenous variables. By this comparison, money and banking appear to matter very much in a number of ways. Costly money and banking push short-term interest rates, $R^B$ and $R^B$, 4 and 5 percentage points lower, respectively, than they would be
otherwise in the steady state.\textsuperscript{14} Money and banking do this by creating a return to collateral and a consequent liquidity services yield that pushes $R^{B}$ far below $R^{T}$, and by making external finance costly in terms of work effort which pushes $R^{IB}$ far below $R^{T}$. Specifically, under the benchmark calibration money and banking create a $LSY^{B}$ of 4% per annum and a real marginal cost of uncollateralized loan production of 5% per annum. Note also, that the calibration implies a ratio of the collateralized to the uncollateralized external finance premium of $(R^{T} - R^{L})/(R^{T} - R^{IB}) = 1 - \alpha = 0.35$.

Money and banking have significant quantitative consequences for macroeconomic aggregates, too. The steady-state capital stock is 5.6% higher in the benchmark calibration than in the counterfactual “highly efficient banking” steady state.\textsuperscript{15} The reason is as follows. The benchmark return $R^{T}$ on the fictitious security that yields no collateral services is invariant to the efficiency of the banking system. Hence, the total net return to capital, which must equal $R^{T}$, is invariant across steady states. At the benchmark calibration $LSY^{K} = kLSY^{B} = 0.2 \times 0.04 = 0.008$, so the marginal product of capital net of depreciation must be 0.8 percentage points lower, which requires a higher $K/n$ ratio. The $K/n$ ratios are 28.7 and 26.5 in the benchmark and highly efficient banking steady states, respectively. The capital stock, $K$, is 5.6% higher in the benchmark calibration in spite of the fact that work in goods production, $n$, is 2.4% lower. The net result is that consumption, $c$, is about 1% lower due to costly money and banking, and total work $m + n$ is about 0.5% lower.

Why exactly do money and banking push the interbank rate $R^{IB}$ down so far according to the benchmark calibration? The reason is that households have an inelastic demand for aggregate bank

\textsuperscript{14} From the perspective of our paper, we would say that the low risk-free rate in Huggett (1993) is due to an implicit broad liquidity services yield on the risk-free asset in his model.

\textsuperscript{15} The extra accumulation of capital in our model is related that in Aiyagari (1994). We would say that the steady-state spread between the rate of time preference and the marginal product of capital in Aiyagari’s model represents an implicit broad liquidity services yield on capital.
deposits in the model which must be funded by borrowing from banks. A no arbitrage condition keeps the uncollateralized loan rate equal to $R^T$. Banks must cover the marginal cost of loan production to willingly accommodate the demand for loans needed to fund the desired deposits. At a zero net interest margin, $R^{IB} = R^T$, a bank would rather lend funds in the interbank market (where the marginal lending cost is zero). An excess supply of interbank credit pushes $R^{IB}$ down to the point where the net interest margin just covers the real marginal cost of loan production. At that point the interbank credit market, the loan market, and the deposit market all clear.

The steady-state implications of money and banking in the benchmark calibration can be appreciated from the perspective of financial economics with reference to the “equity premium puzzle” which asks why real stock returns average over 6 percentage points per annum above real short-term interest rates.\(^{16}\) We saw above that the inclusion of money and banking implies a liquidity services yield that pushes $R^B$ about 4 percentage points below $R^T$, and competitive banks that push $R^{IB}$ over 5 percentage points below $R^T$. However, according to the benchmark calibration the liquidity services yield on capital only pushes the net marginal product of capital 0.8 percentage points below $R^T$ in the steady state. Thus, according to the benchmark calibration, money and banking push real interest rates down relative to the return to capital, which we identify with equity, by a net of 4.2 percentage points for $R^{IB}$ and 3.2 percentage points for $R^B$. Thus, it would appear that considerations of the kind related to money and banking identified in our model make a quantitatively significant contribution to resolving the equity premium puzzle.\(^{17}\)

In passing, it is worth noting that in the finance literature the theoretical interest rate defined by the pricing function (16), $R^T$, is generally identified empirically with a short-term interest rate such as $R^{IB}$ or $R^B$, all being referred to as the “risk-free rate.” The rates may be risk free in theory

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\(^{16}\) See, e.g., Campbell (1999, p. 1234).

\(^{17}\) See Bansal and Coleman (1996) for an asset-pricing model that also distinguishes between explicit pecuniary and implicit service yields in a study of the equity premium.
and in practice, at least if inflation is stabilized sufficiently. However, our model suggests that in the presence of considerations giving rise to money and banking, this identification is quantitatively inappropriate and potentially highly misleading.

We now move on to report two experiments with the benchmark calibration to gauge the sensitivity of steady state outcomes to the debt to GDP ratio and to the velocity of aggregate bank deposits. Among other things, such experiments indicate the magnitude of the variation that a central bank can expect in the average real interbank interest rate, i.e., “real neutral interbank interest rate,” that is consistent with the economy growing at trend with full employment and with the price level stabilized on a path consistent with the central bank’s inflation target.

Table 2 shows two counterfactual steady-state calibrations for the ratio of public debt to consumption, one in which “boc” is 50% higher than in the benchmark calibration and another in which boc is 50% lower. The steady state is sensitive to the stock of public debt relative to consumption because boc controls the weighted value of collateral in the economy. Macroeconomic aggregates such as consumption, employment in goods production, and the real wage are relatively invariant to boc. Some crowding out (and crowding in) of the capital stock is evident, K falls by about 1.6% in response to a 50% increase in boc, and K rises by 2.4% with respect to a 50% fall in boc. Employment in banking is relatively sensitive to boc, falling by about 1/6 with a 50% increase in boc, and rising by around 1/3 with a 50% fall in boc.

Except for the fictitious rate $R^T$ which is, of course, invariant with respect to the ratio of public debt to consumption in the steady state, interest rates exhibit substantial sensitivity to boc. The 50% increase in boc raises $R^{IB}$ by 0.8 percentage points per year. And the 50% reduction in boc causes $R^{IB}$ to fall by 1.2 percentage points per year, driving the real interbank interest rate about 30 basis points per year below zero. These interest rate effects reflect changes in the real marginal

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18 We interpret this experiment as if the steady-state rate of inflation is high enough to keep $R^{IB}$ above the zero bound.
cost of loan production associated with the effect of boc on the weighted value of collateral. The government bond rate $R^B$ responds similarly, rising by 1.2 percentage points per year with the 50% increase in boc and falling by a similar magnitude with boc. The consequences for $R^B$ reflect the fact that liquidity service yields on assets move in the opposite direction from the weighted value of collateral. The uncollateralized external finance premium, UEFP, reflects this, moving inversely by about 40 basis points per year with the 50% displacement of boc relative to the benchmark. On the whole, our experiments with the benchmark calibration suggest that the “neutral” real interbank interest rate consistent with stable inflation varies in a quantitatively significant way with the ratio of public debt to GDP.

Table 3 shows two counterfactual steady-state calibrations in which velocity of aggregate bank deposits “$V$” is 10% higher and lower, respectively, than in the benchmark calibration. The steady state is sensitive to V because velocity governs the demand for bank deposits and thereby the induced demand for bank loans to fund deposits. As was the case with respect to “boc,” the macroeconomic aggregates c, n, and w are relatively insensitive to velocity. The capital stock is somewhat sensitive to velocity, falling by 1.3% with a 10% rise in V, and rising by 1.7% with 10% lower V. The sensitivity of the capital stock is due to the fact that the liquidity service yields on bonds and capital move inversely with respect to velocity, since lower velocity raises the demand for collateral services in loan production. Hence, the stock of capital must adjust so that the sum of the net marginal product of capital and the liquidity services yield on capital sum to $R^T$, which is invariant to velocity in the steady state.

The liquidity service yield effect on bonds causes $R^B$ to vary directly by 1.2 percentage points per year with a 10% displacement of V relative to the benchmark calibration. The uncollateralized external finance premium moves inversely with V by about 40 basis points per year for a 10% displacement of V relative to the benchmark. The effect on UEFP is connected with the
inverse movement of employment, m, in banking which, in turn, reflects the inverse relationship between velocity and the demand for banking services.

Finally, \( R^{IB} \) exhibits substantial sensitivity to velocity that reflects the inverse relationship between velocity and the real marginal cost of loan production. \( R^{IB} \) rises by 0.8 percentage points per year when \( V \) is 10% above the benchmark and \( R^{IB} \) falls by 1 percentage point when \( V \) is 10% below the benchmark. The real interbank rate is about 10 basis points per annum below zero in the latter case. On the whole, our experiments with the benchmark calibration suggest that the “neutral” real interbank interest rate consistent with stable inflation varies in a quantitatively significant way with the velocity of aggregate bank deposits.

7. Linearization

We now move on to the dynamic issues mentioned in the introduction. In that regard we need to complete the specification of our dynamic model, linearize it in a suitable manner, and calibrate the new components. The first step is to recognize the presence, irrelevant in steady-state analysis, of sluggish price adjustments. For that purpose we adopt the most nearly standard of all price-adjustment formulations, the discrete-time version of the Calvo (1983) mechanism. We recognize that there are grounds for objection to this specification, but as our model is strongly unorthodox in other ways, we consider it best to stay close to the mainstream with this component. Accordingly, we adopt the following price adjustment equation, applicable at the household level and also to the general price level for goods in a symmetric equilibrium:

\[
\Delta p_t = \beta E_t \Delta p_{t+1} + \kappa m c_t + u_t \quad \kappa > 0.
\]

Here \( p_t = \log P_t = \log P_t^A \) so \( \Delta p_t \) represents the inflation rate, while \( m c_t \) is the real marginal cost of goods production. With sticky prices, relation (13) above is replaced by

\[
mc_t = \xi_t / \lambda_t,
\]

so marginal cost is not constant in this case. We have added two equations and only one variable
(not counting \( p_t \) or its defining equation) so some relationship from Section 2 has to be eliminated. It is, of course, the labor demand function (10), which represents behavior under full price flexibility.

The other major addition to the model of Section 2 is a specification of monetary policy behavior by the central bank. In the policy experiments below we will use two alternative types of policy rules. The first is a rule for setting the interbank interest rate, \( R_t^{IB} \). Specifically, we adopt a rule of the general type made famous by Taylor (1993), allowing (as in most recent work) inclusion of a lagged value of \( R_t^{IB} \) to reflect the prevalence of interest-rate smoothing by many central banks:

\[
R_t^{IB} = (1 - \mu_1)[\mu_0 + (1 + \mu_1)\Delta p_t + \mu_2mct_t] + \mu_3R_{t-1}^{IB} + e_t.
\]

Here we have written the rule in a fashion that implicitly sets the central bank’s target inflation rate, say \( \pi^* \), equal to zero. With a Cobb-Douglass production function, \( mct \) serves as a measure of the output gap used by Taylor (1993).

While (34) expresses monetary policy behavior in a rather realistic way, it will be useful for diagnostic purposes also to consider some policy experiments based on an alternative type of rule that depicts the central bank as controlling the growth rate of the stock of high-powered (base) money, \( H_t \). Specifically, we will also report results based on the following monetary policy rule, in which \( \Delta \hat{H}_t \) is the growth rate of \( H_t \):

\[
\Delta \hat{H}_t = \rho^H \Delta \hat{H}_{t-1} + e_t^H.
\]

Here and in what follows, we use symbols with a “hat” to represent fractional deviations from steady-state values. Also, in (34’) we have \( 0 < |\rho^H| < 1 \) while \( e_t^H \) represents the random, unsystematic component of policy behavior. In our model, \( \hat{H}_t \) is related to consumption and the price level by equations (4) and (5); the implied linearized relationship is

\[
\hat{H}_t = \hat{c}_t + \hat{P}_t.
\]
Equations (32), (33), (34) [or (34')] and (35) are combined with linearized versions of equations (4)-(15). Because we are using a symmetric equilibrium condition, the number of equations needed is smaller than in Section 2. Essential relationships are as follows:

\[
\hat{\lambda}_t + \hat{w}_t = \left(\frac{n}{1-n-m}\right)\hat{n}_t + \left(\frac{m}{1-n-m}\right)\hat{m}_t
\]

\[
\hat{w}_t + \hat{m}_t + \left(\frac{c(1-\alpha)}{mw}\right)\hat{c}_t + \left(1 + \frac{c(1-\alpha)}{mw}\right)\hat{\lambda}_t = 0
\]

\[
\hat{\Omega}_t = \left(\frac{kK/c}{b + (kK/c)}\right)(\hat{c}_t - \hat{q}_t - a\hat{3}_t) - \left(\frac{b}{b + (kK/c)}\right)\hat{b}_t , \text{ with } b = B/(P(1+R^B)c)
\]

\[
\hat{c}_t = \left(1 + \frac{\delta K}{c}\right)(1-\eta)(\hat{n}_t + a\hat{1}_t) - \left(\frac{\delta K}{c}\right)\hat{q}_t
\]

\[
\hat{c}_t = (1-\alpha)(a\hat{2}_t + \hat{m}_t) + \left(\frac{abc}{bc+kK(1+\gamma)}\right)(\hat{b}_t + \hat{c}_t) + \left(\frac{\alpha kK(1+\gamma)}{bc+kK(1+\gamma)}\right)(a\hat{3}_t + \hat{q}_t)
\]

\[
m\hat{c}_t = \hat{n}_t + \hat{w}_t - \hat{c}_t
\]

\[
\Delta\hat{p}_t = \hat{p}_t - \hat{p}_{t-1}
\]

\[
\left(\frac{k\Omega\phi}{c\lambda}\right)\hat{c}_t =
\]

\[
k\Omega\left[\frac{\phi}{c\lambda} - 1\right](\hat{\Omega}_t + a\hat{3}_t) + \left[k\Omega\left(\frac{\phi}{c\lambda} - 1\right)\right]\hat{q}_t + \frac{\beta}{1+\gamma}[1-\delta + \eta mc(n/K)^{1-\eta}]E_t\hat{\lambda}_{t+1} +
\]

\[
\frac{\beta}{1+\gamma}(1-\delta)E_t\hat{q}_{t+1} + \left[\frac{\beta \eta mc}{1+\gamma}(n/K)^{1-\eta}\right]E_t(m\hat{c}_{t+1} + (1-\eta)\hat{n}_{t+1} + a1_{t+1}) - \left[\frac{k\Omega\phi}{c\lambda} + \frac{\beta}{1+\gamma}(1-\delta + \eta mc(n/K)^{1-\eta})\right]\hat{\lambda}_t
\]

The last of these, which is the linearized version of the first order condition with respect to the capital stock, is obviously a complex condition. Be that as it may, inspection of equations (32), (33), and (35)-(43) shows that the endogenous variables that appear in this system are

\[\Delta\hat{p}_t, m\hat{c}_t, \hat{\Omega}_t, \hat{\lambda}_t, \hat{c}_t, \hat{q}_t, \hat{n}_t, \hat{m}_t, \hat{c}_t, \hat{q}_t, \text{ and } \hat{p}_t.\] Thus, with policy determination of the paths of
\(\hat{H}_i\) and \(\hat{b}_i\), we have the proper number of relationships for this essential set of variables.\(^19\) Once their solutions are found, it is straightforward to calculate the interest rates \(R_i^T, R_i^B, R_i^{IB},\) and \(R_i^L\) using linearized versions of the relationships (16), (17), (21), and (22) given in Section 3.

8. Dynamic Analysis

We study dynamics using the log-linearized model calibrated at the steady state, as amended in Section 7 to include Calvo-style staggered pricing and a policy rule, to explore how much money and banking matter quantitatively for an otherwise standard “new neoclassical synthesis” (aka, “new Keynesian”) model of monetary policy. The dynamic consequences of money and banking for monetary policy are illustrated in two ways.

First, we demonstrate the presence of two distinct channels by which money and banking affect the transmission of monetary policy. We identify a banking “attenuator” channel that tends to blunt the effect of monetary policy actions. However, there is also a banking “accelerator” channel in the model resembling the familiar financial accelerator with the potential to amplify the effect of monetary policy actions. We discuss the mechanics and implications of the two potentially conflicting channels of monetary transmission for the behavior of the economy and the effectiveness of monetary policy.

Second, we illustrate how much a central bank could be misled quantitatively in its interbank rate policy actions by ignoring money and banking altogether. To this end, we investigate the extent to which a central bank could misjudge its interest rate policy response to a generic goods productivity shock by not taking money and banking into account. Then we consider two shocks emanating from the banking sector itself (which have no counterpart in the benchmark new

\(^{19}\) Of course, the dynamic solution requires proper attention to side conditions, etc. In all the results reported below, the utilized solution is the only stable solution.
synthesis model)—a shock to loan monitoring productivity and a shock to effective collateral meant to capture an effect of widespread financial distress.

The banking attenuator and accelerator channels of monetary policy transmission are highlighted, respectively, in Figures 1 and 2. Each figure reports impulse responses of key endogenous variables to a 1 percentage point shock to policy rule (34’) assumed to generate the growth of high-powered money. The experiments depicted in the two figures differ only in the assumed autoregressive (AR) coefficient of the money growth rule. Figure 1 reports outcomes for a modest degree of money growth persistence, AR = 0.6. Money growth in Figure 2 is highly persistent with AR = 0.99.

In Figure 1, the modest degree of money growth persistence stimulates a 0.6 percentage point increase in consumption on impact that dies away relatively quickly as firms adjust prices to neutralize the cumulative money growth on the markup and, thereby, on other real variables. The aggregate demand stimulated by expansionary monetary policy is accommodated by an expansion of hours worked in goods production and an elevated real wage. There is also a relatively large increase in the (consumption) price of capital, q, due to the elevated capital-labor ratio and the elevated marginal product of capital.

The banking “attenuator” is evident in Figure 1 in the following sense. According to the “broad liquidity constraint” (25), the increase in consumption generates a proportionate increase in the demand for bank deposits that must be accommodated by a commensurate increase in loan production. The increase in the value of the capital stock associated with the rise in q provides additional collateral value that helps produce more loans. However, according to the simulation the banking system also expands loan production effort, m, to accommodate the increased demand for deposits and loans.

Figure 1 shows that the net effect, according to (23), of m, w, and c is to raise the
uncollateralized external finance premium by around 0.9 percentage points. Thus, the $R^T - R^{IB}$ interest rate spread rises by 0.9 percentage points. Both nominal interest rates rise initially by 0.5 percentage points reflecting inflation expectations. $R^T$ first rises and then falls relative to expected inflation due to movements in the pure real rate of interest associated with the expected change in $\lambda$ according to (16). Consequently, $R^{IB}$ itself falls initially by only 0.5 percentage points on net.\(^{20}\)

In this case the external finance premium is procyclical—it rises in response to an expansionary monetary policy shock with the real marginal cost of loans. If the model were to allow for investment financed externally at the margin, the procyclical external finance premium would “attenuate” the effect of monetary stimulus on investment. As it stands, an attenuation is present in the model because effort is drawn into banking with the expansion of goods production. The procyclical effect on the external finance premium reflects the fact that monetary stimulus on consumption is accompanied by an increased demand for banking services to accommodate the demand for bank deposits. The external finance premium is procyclical, rather than countercyclical as in the work of Bernanke, Gertler and Gilchrist (1999) among others that neglects the direct relationship between consumption and the demand for bank deposits through the “demand for money.”

Figure 2 reports the results of a highly-persistent money growth shock that generates a highly persistent 1 percent rate of inflation, nearly twice the peak response of inflation in Figure 1. The higher rate of inflation in this case neutralizes much of the impact on consumption, which now rises by only 0.1 percentage points before it begins to fall. Hours worked in goods production and the real wage peak after a lag at only 0.2 percentage points above the steady state, respectively. However, the q price of capital still rises by around 0.4 percentage points and remains persistently elevated. The increased value of capital collateral relative to consumption allows broad liquidity

\(^{20}\) The reported decline in $R^{IB}$ is relative to the steady-state. We interpret this simulation under the condition that the steady-state rate of inflation is sufficiently positive that $R^{IB}$ remains above the zero bound.
constraint (25) to be satisfied with a substantial decline of nearly 0.5 percentage points of hours worked in banking.

Figure 2 shows that the net effect, according to (23), of m, w, and c is to lower the external finance premium persistently by as much 0.4 percentage points, yielding a commensurate reduction in the $R^T - R^{IB}$ interest rate spread. $R^{IB}$ rises by 1.4 percentage points on impact reflecting the sum of inflation expectations and the countercyclical external finance premium. The banking “accelerator” effect on monetary stimulus in this case is reflected in the fact that the banking sector requires less work effort to satisfy the increased demand for bank deposits, so that the real marginal cost of loan production falls. The present discounted value of the persistently elevated future marginal product of capital keeps the q price of capital high enough on impact to overcome the presence of the banking attenuator and support a net banking accelerator.

The mechanics of the banking accelerator are much the same as in Bernanke, Gertler, and Gilchrist (1999) and others where a countercyclical external finance premium amplifies and propagates monetary policy shocks through a positive feedback loop that involves the monetary stimulus, asset returns, the q price of capital, and the consumption value of loan collateral. The difference is that in our money and banking model the accelerator competes with an attenuator. As it is currently specified, our model is unduly favorable quantitatively to the attenuator because all output is consumed and the velocity of bank deposits with respect to consumption is assumed constant. Realistic modifications of the model—allowing output to be invested and the velocity of deposits to vary—could shift the balance considerably. The demand for deposits could be related more closely and realistically to consumption than to investment, and the consumption velocity of money could be made interest sensitive. In that case, consumption would be smoothed relative to investment and output, and velocity would be procyclical. Both outcomes would strengthen the accelerator relative to the attenuator.
We now proceed to ask: How much could a central bank that ignores money and banking misjudge the proper interest rate policy action to stabilize inflation in response to a goods productivity shock? We answer the question by comparing the responses of the real natural interbank rate of interest to a goods productivity shock in the model with and without money and banking. We do this by utilizing interest rate rule (34) with \( \mu_1 = 50, \mu_2 = 0, \) and \( \mu_3 = 0. \) To the approximation reported by Matlab, this specification approximately stabilizes inflation, which according to price adjustment equation (32) with the \( u_t \) shock turned off, stabilizes real marginal cost “mc” of goods production at its steady-state value. Thus, this specification of the interest rate rule essentially makes nominal interest rates behave like real natural interest rates consistent with the underlying imperfectly competitive real business cycle model with flexible prices and a constant profit-maximizing markup.

We shock the linearized model with a 1 percentage point, highly-persistent first-order autoregressive goods productivity process with AR coefficient 0.99. The impulse responses shown in Figure 3 indicate that the following endogenous variables are nearly invariant to the shock: \( n, dp, R^T, \) and \( R^B. \) The model has log utility and non-storable output, so we would expect hours worked in goods production to be relatively insensitive to the highly-persistent shock to the level of productivity in goods production. The inflation rate is essentially stabilized, as is the benchmark nominal interest rate \( R^T \) according to (16), because expected inflation and the expected change in \( \lambda \) are approximately zero. The relative stability of the real benchmark natural rate of interest rate reflects the fact that desired household saving is relatively insensitive to the highly persistent shock to the level of productivity.

The striking contrast is that \( c, w, q, m, \) and the external finance premium all rise persistently by 0.6 or 0.7 percentage points. Why? With hours in goods production relatively stable, increased productivity shows up as a jump in consumption approximately equal to labor’s share. The jump in
the wage reflects the increase in labor productivity. And the q price of capital reflects the highly persistent increase in the marginal product of capital. The increase in banking hours reflects the increased demand for deposits associated with the rise in consumption according to the broad liquidity constraint. The small decrease in hours worked in goods production partly reflects the increase in banking hours. As a result, on net according to (21) the uncollateralized external finance premium jumps by 0.7 percentage points.

These results reveal a striking quantitative effect of the presence of money and banking in the model. A central bank unaware of the effect of money and banking on interest rate spreads in general and on the interbank rate in particular would expect all real natural interest rates to behave like $R^T$. Hence, a central bank determined to make its interbank rate policy instrument shadow movements in the real natural rate so as to stabilize inflation perfectly would move its interbank rate relatively little in response to a highly persistent productivity shock. Yet, the model with money and banking requires the central bank to cut its interbank rate by around 0.7 percentage points to stabilize inflation fully in response to such a shock. That amounts to a 2.8 annualized percentage point misjudgment of the desired interbank rate policy action.

What if the central bank followed the original Taylor Rule? To find out, we compute impulse responses to a 1 percentage point positive goods productivity shock with AR coefficient 0.99, but with interest rate rule (34) calibrated as $\mu_1 = 1.5$, $\mu_2 = 0.5$, and $\mu_3 = 0$. The striking result reported in Figure 4 is that the Taylor Rule creates a persistent 5 percent deflation! The deflation is associated with a persistent decline in real marginal cost (an elevated markup) in goods production. Expected deflation pushes all three nominal interest rates lower. Hours in goods production, consumption, the wage, and the price of capital are all marginally lower than in Figure 3. The key banking variables, monitoring effort and the uncollateralized external finance premium, are both persistently higher under the Taylor Rule.
Money and banking make the Taylor Rule deflationary. If money and banking were free or absent from the model altogether, then there would be a single interest rate in the model, $R^T$. And that real interest rate would not move much in response to the highly persistent goods productivity shock because households would not wish to save much of it. The permanent income theory of consumption says that aggregate demand would increase to absorb most of the increase in aggregate supply without much variation in the real interest rate. Hence, the output gap would be small, and deflation could be averted with a relatively small cut in the interest rate.

According to our calibration, however, the real interbank rate must fall by a lot relative to the real benchmark rate, $R^T$. However, the Taylor Rule says that the real interbank rate cannot fall if inflation and the output gap are stabilized. Hence, in dynamic equilibrium the economy must settle at a point where there is just enough of an output gap (elevated markup) and deflation for the Taylor Rule to induce a cut in the real interbank rate to satisfy equilibrium in the banking sector. Thus, our model of money and banking creates a quantitatively significant problem for the Taylor Rule.

That said, the potential for large mistaken interbank rate policy actions might be reduced if the model were modified to allow banks to acquire government bonds as assets. As currently specified, we have not allowed a bank to borrow in the interbank market in order to buy and hold government bonds outright. Since a bank would incur little loan production costs on either side of this arbitrage, a bank would profit by utilizing funds borrowed in the interbank market to buy government bills as long as $R^B$ exceeded $R^{IB}$.

Arbitrage transactions would have two important effects on interest rates. On one hand, given the outstanding stock of government bonds, the acquisition of government bonds by banks would reduce collateral against which households could borrow. On the other hand, the purchase of government bonds by banks would reduce the volume of loans needed to accommodate a given
demand for deposits. Consequently, arbitrage would push up the liquidity service yield on assets as collateral, and it would push down the external finance premium. The higher liquidity service yield would lower $R^B$ relative to $R^T$, and the lower external finance premium would raise $R^{IB}$ relative to $R^T$. Banks would purchase government bonds up to the point where the short-term government bill rate is arbitrated into equality with the interbank rate. In this way, allowing banks to buy government bonds outright in the model would increase the sensitivity of $R^B$ and possibly $R^T$ to the productivity shock, but it would mitigate the fall in $R^{IB}$ needed to stabilize inflation in response to the productivity shock.

Our remaining examples of the quantitative consequences of money and banking involve two shocks that emanate from the banking sector itself which have no counterpart in the benchmark new synthesis model: a shock to the productivity of monitoring effort, and a shock to effective collateral.

The detrended productivity of monitoring effort in banking is represented by $a_{2t}$ in the linearized broad liquidity constraint (40), and is determined exogenously by a first-order autoregressive process. We assume an AR coefficient of 0.99 for the banking productivity shock to capture the persistent nature of surprising improvements or disappointments in technological progress in banking relative to trend. And we assume the same specification for the interest rate rule employed above to stabilize inflation fully.

Figure 5 displays impulse responses to the banking productivity shock. As expected, the inflation rate is stabilized reasonably well, as are $R^T$ and $R^{IB}$. Hours worked in banking fall by 0.85 percentage points in response to improved productivity in banking. In part, this is because there is little change in consumption and therefore, little change in the demand for deposits and loans. There is little change in the q value of collateral, and therefore little effect on the supply of loans for given hours worked in banking. According to (21) the fall in $m$, together with the negligible change
in the wage and the small increase in consumption, produces a 0.9 percentage point fall in the uncollateralized external finance premium, which is reflected in an equivalent rise in $R^{IB}$. This means that a 1 percent improvement in banking technology requires the central bank to raise its real interbank rate policy instrument by 3.6 percentage points at an annual rate to stabilize inflation.

Here again money and banking potentially matter a great deal quantitatively for interest rate policy. A central bank unaware of the consequences of money and banking that wished to shadow the natural real interbank rate to stabilize inflation would miss completely the need to raise its interbank rate instrument to do so. The central bank would be surprised at the resulting strength of economic activity and inflationary pressure. Impulse responses not reported here to save space show that a central bank following the original Taylor Rule in this case would create inflation persistently in excess of 0.5 percent.

That said, there are reasons why shocks to banking technology need not precipitate such large movements in the natural interbank rate. First, the natural interbank rate is invariant to productivity shocks when banking and goods productivity trend upward together as in the steady state. Second, the variance of detrended banking productivity shocks that move independently of goods productivity shocks might be relatively small, with modest implications for the natural interbank rate. Third, if banks were allowed to hold government securities outright in the model, then arbitrage between $R^{IB}$ and $R^{B}$ would induce banks to sell government securities in response to the positive banking productivity shock, which would mitigate the rise in the natural interbank rate.

For our last experiment we consider a shock meant to capture the consequences of financial distress that weakens a borrower’s property rights to his capital and makes capital less productive in securing and producing loans. We model financial distress as a shock to “k,” the coefficient which transforms a given value of the capital stock into effective collateral in loan production function (6). The shock to effective collateral shows up as a shock to $a_{3t}$ in the linearized equations (38) and (40).
We assume that $a_3$ is generated as an exogenous first-order autoregressive process with an AR coefficient of 0.9 to capture modest persistence associated with resolving financial distress. Again, we specify the interest rate rule as above to stabilize inflation fully in order to illustrate the effect of financial distress on the natural interbank rate of interest.

Figure 6 displays impulse responses to a 1 percentage point negative shock to effective collateral. The inflation rate is stabilized reasonably well, as are $R^T$ and $R^B$. Hours worked in banking rise by 1.2 percentage points in response to the decline in effective collateral. This is because there is little change in $c$ and $q$, and so little change in the demand for deposits and in the supply of loans given hours worked in banking. According to (21) the rise in $m$, together with the small changes in $c$ and $w$, produce a 1.2 percentage point rise in the uncollateralized external finance premium, reflected in an equivalent fall in the natural real interbank interest rate.

Once more, developments relating to money and banking potentially matter a great deal for interest rate policy. In this case, it is implausible that a central bank would remain unaware of the money and banking shock. Evidence of financial distress would be directly observable and reported in the media. Central banks generally recognize the need to cut the interbank rate in response to widespread financial distress. However, given the absence of money and banking in models of monetary policy, there is no way to judge how much to cut the interbank interest rate.

The simulation indicates that a 1 percentage point decline in effective collateral would require an initial 4.8 annualized percentage point cut in the interbank interest rate followed by a gradual return to neutral. This sensitivity of the natural interbank rate to widespread financial distress seems large in the sense that central banks rarely make such immediate, deep cuts in the interbank rate, even in response to a financial crisis.

To get a more realistic idea of the effect of financial distress, we consider a more realistic interest rate rule in Figure 7. Figure 7 displays impulse responses to the negative collateral shock
when monetary policy is governed by a specification of interest rate policy (34) reflective of a benchmark Taylor Rule with a lagged interest rate: $\mu_1 = 0.5$, $\mu_2 = 0.5$, and $\mu_3 = 0.8$. In this case, interest rate smoothing delays the cut in the interbank rate, which reaches a trough after about one year only 2.5 percentage points below neutral. Realistic interest rate policy produces more realistic macroeconomic outcomes: there is a recession with a 2 percent contraction in hours worked and a 2 percent per annum deflation. In this sense, the sensitivity of the natural real interbank rate to the effective collateral shock evident in the simulation may not be implausibly large.

In fact, financial distress is likely to be accompanied by a flight to safety resulting in an increased demand for bank deposits relative to consumption. The decline in velocity would cause an even larger fall in the natural interbank rate than in Figure 6. On the other hand, if banks were allowed to hold government securities outright in the model, then arbitrage between $R^{\text{RB}}$ and $R^B$ would induce banks to buy government securities in response to the financial distress. Such arbitrage would bring $R^B$ and $R^T$ down, cushion the decline in $R^{\text{IB}}$, and improve the capacity of the model economy to respond to financial distress.

9. Conclusions

We conclude with a brief summary of the paper’s scope and findings. Our objective was to reconsider the role of money and banking in monetary policy analysis by means of an analytical framework that includes both a banking sector and transaction-facilitating money in an optimizing model that is otherwise of the standard new neoclassical synthesis type. The addition of the banking sector leads to several unusual features, including a number of distinct interest rates. The model is implemented quantitatively, based on a calibration that attempts (within specification constraints) to be realistic for an economy such as that of the United States.

Results obtained are of two types, pertaining to steady-state and dynamic properties of the model. In the former case we are reasonably successful in providing an endogenous explanation for
substantial steady-state differentials between the short-term interbank interest rate, often used as the central bank’s policy instrument, and the following: (i) the collateralized loan rate, (ii) the uncollateralized loan rate, (iii) the T-bill rate, (iv) the net marginal product of capital, and (v) a benchmark risk-free rate. One steady-state experiment involves a counterfactual calibration that makes banking services almost costless to produce; this scenario results in interbank and T-bill rates that are 4 percentage points per annum higher than with our baseline calibration, thereby indicating a major quantitative effect of a banking sector. Effects on the steady-state capital stock are also sizeable; with costless banking services, the capital stock is over 5 percent lower than in the baseline calibration. Among other results, in the baseline calibration we find a differential of over 3 percentage points per annum between short-term interest rates and the return to capital, a magnitude that contributes significantly to resolving the famous “equity premium puzzle.” Finally, we report experiments with the baseline calibration that indicate a sizable sensitivity of the steady-state “neutral” interbank rate to the debt to GDP ratio and to the velocity of aggregate bank deposits.

Dynamic results are based on impulse response functions implied by a version of the model that is log-linearized around the steady state. Here we demonstrate the quantitative significance of a new component of monetary policy transmission, a “banking attenuator,” that works in the opposite direction from the “financial accelerator” of Bernanke, Gertler, and Gilchrist (1999)—which itself is also present in the model, nevertheless.

One of the more significant findings in this dynamic context is that, according to the model, a central bank that based its rate-setting policy on analysis that failed to recognize the distinction between the interbank rate and the benchmark risk-free rate would miss its appropriate settings by as much as 4 percentage points. For instance, we show that a central bank that utilized an interbank interest rate instrument with parameters chosen as in the original Taylor (1993) rule would produce a persistent 2 percent per annum deflation in response to a 1 percent, highly-persistent positive
shock to productivity in goods production.

Finally, we demonstrate the quantitative consequences of two shocks emanating from the banking sector itself: a shock to banking productivity and a shock to effective collateral reflecting financial distress. For instance, we show that the central bank should cut the real interbank rate initially by nearly 5 percentage points per annum to fully stabilize inflation and output against a moderately persistent, 1 percent decline in effective collateral. Moreover, we show that a more realistic policy response reflecting instead a benchmark Taylor Rule with a lagged interest rate fails to offset the contractionary consequences of financial distress, instead producing a recession with 2 percent per annum deflation and a 2 percent contraction in employment. In short, in our calibrated model, the effects of money and banking are quantitatively of major importance.
**Table 1: Steady-State Calibrations**

**Benchmark Calibration**  
(quarterly data; zero inflation)

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**Highly Efficient Banking**  
(Same calibration as above except $F = 90.$)

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**Table 2: Sensitivity of Steady State to Ratio of Public Debt to GDP**

**Public Debt to GDP in the Steady State**  
(50% higher and lower than the 0.56 benchmark calibration)

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Table 3: Sensitivity of Steady State to Velocity

Velocity in the Steady State
(10% higher and lower than the 0.31 benchmark calibration)

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Figure 1: Responses to unit shock to edh

Figure 2: Responses to unit shock to edh
Figure 3: Responses to unit shock to $a_1$

Figure 4: Responses to unit shock to $a_1$
Figure 5: Responses to unit shock to $a_2$

Figure 6: Responses to unit shock to $-a_3$
Figure 7: Responses to unit shock to -a3
References


McCallum, B.T., 2002. Recent developments in monetary policy analysis: the roles of theory and